FRACTIONAL ORNSTEIN-UHLENBECK SIGNAL PROCESSING

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Abstract

In this paper we consider a problem of signal processing where the signal is expressed by a fractional Ornstein-Uhlenbeck process in general form. An explicit form of the signal is derived from a fractional Langevin equation. A method of L^2 -approximation is used to find the approximate estimate for the state of the fractional signal and the convergence to the optimal estimate is established.

1 Introduction

It is known that the Ornstein-Uhlenbeck plays a crucial role in telecommunication as an only stationary Gaussian Markov signal with white noise. But a Gaussian non-Markovian signal is also important in some context where the signal leaves a long time influence upon its behavior. A good candidate for expressing this signal property is a fractional Brownian noise. In this paper we consider a problem of signal processing where the signal is a fractional Ornstein-Uhlenbeck by introducing an approximation approach.

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1.1 Fractional Brownian motion

A fractional Brownian motion of Mandelbrot form is a centerred Gaussian process $(W_t^H, t \ge 0)$ with covariance function R(s, t) given by

$$R(s,t) = E(W_s^H W_t^H) = \frac{1}{2}(s^{2H} + t^{2H} + |t-s|^{2H}),$$
(1.1)

where H is a parameter called Hurst index, 0 < H < 1. In the case where $H = \frac{1}{2}$, W_t^H becomes a usual standard Brownian motion. The process W_t^H can be decomposed as

$$W_t^H = C_H (U_t + B_t^H), (1.2)$$

where U_t is a stochastic process with absolutely continuous trajectory and C_H is a constant depending only on H, $B_t^H = \int_0^t (t-s)^{\alpha} dW_s$ with $\alpha = H - 1/2$. We know that W_t^H is a process of long memory with $H \neq \frac{1}{2}$. In (1.2) this property focuses at the second term B_t^H and by this reason, B_t^H is called a fractional Brownian motion of Liouville form. In this paper we consider fractional noise associated with B_t^H . The problem is how to get the optimal state estimation for a fractional signal that is a general fractional Ornstein-Uhlenbeck process $(X_t, t \geq 0)$ satisfying the following equation

$$dX_t = (a(t)X_t + b(t))dt + \sigma dB_t^H, \qquad (1.3)$$

where H > 1/2, from an observation Y_t given by

$$dY_t = h(X_t)dt + dV_t, (1.4)$$

where V_t is a standard Brownian motion independent of B_t^H , $h_t = h(X_t)$ is a process such that

$$E \int_0^t h_s^2 ds < \infty$$
, for every $t \ge 0$.

1.2 Approximation approach

The fractional Brownian motion B_t^H is not a semimartingale, so a fractional signal driven by B_t^H as X_t in (1.3) cannot be solved by the traditional Ito calculus.

An L^2 -approximation approach has been introduced in [2] where a process $B_t^{H,\epsilon}$ is considered instead of B_t^H :

$$B_t^{H,\epsilon} = \int_0^t (t-s+\epsilon)^{\alpha} dW_s, \ \alpha = H - \frac{1}{2}.$$
 (1.5)

A calculation says to us that $B_t^{H,\epsilon}$ is in fact a semimartingale.

$$dB_t^{H,\epsilon} = \alpha \varphi_s^{\epsilon} dt + \epsilon^{\alpha} dW_t, \qquad (1.6)$$

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where

$$\varphi_t^{\epsilon} = \int_0^t (t - s + \epsilon)^{\alpha - 1} dW_s, \ \alpha = H - \frac{1}{2}.$$

And as shown in [2] we have the following fundamental result on L^2 -convergence of semimartingales $B_t^{H,\epsilon}$. **Result:** $B_t^{H,\epsilon}$ converges to B_t in $L^2(\Omega)$ when $\epsilon \to 0$ and we have

$$\sup_{0 \le t \le T} \|B_t^{H,\epsilon} - B_t\|_{L^2} \le K(\alpha) e^{1/2 + \alpha}, \tag{1.7}$$

where $K(\alpha)$ is a constant depending only on $\alpha = H - 1/2$.

Moreover a new approach to stochastic integration and stochastic differential equations driven by B_t^H is given in [3] (refer also to [3]-[9]).

$\mathbf{2}$ General fractional Ornstein-Uhlenbeck signal

2.1Approximate signal equation

Consider again the equation

$$dX_t = (a(t)X_t + b(t))dt + \sigma dB_t^H, H > 1/2,$$
(2.1)

where $0 \le t \le T$, coefficients a(t) and b(t) are deterministic continuous function on [0, T].

It is a generalization of fractional stochastic Langevin equation studied in [6] and [7], where our L^2 -approximation method has been applied to find its solution. As shown in [10] the solution of (2.1) is a L^1 -limit of that of an approximate equation. Now we prove that it is also a L^2 -limit. By replacing B_t^H by $B_t^{H,\epsilon}$ we obtain the approximate equation for the signal

 X_t as follows

$$dX_t^{\epsilon} = (a(t)X_t^{\epsilon} + b(t))dt + \sigma dB_t^{H,\epsilon}, \qquad (2.2)$$

where $0 \le t \le T$, H > 1/2.

2.2Approximate equation

Equation (2.2) can be rewritten as follows

$$dX_t^{\epsilon} = (a(t)X_t^{\epsilon} + b(t) + \alpha\varphi_t^{\epsilon})dt + \sigma\epsilon^{\alpha}dW_t.$$
(2.3)

A method of equation splitting introduced by us in [6, 7] can be applied to (2.3). We can write

$$X_t^{\epsilon} = X_1^{\epsilon}(t) + X_2^{\epsilon}(t), 0 \le t \le T,$$

$$(2.4)$$

where

$$dX_1^{\epsilon}(t) = a(t)X_1^{\epsilon}(t)dt + \sigma\epsilon^{\alpha}dW_t \tag{2.5}$$

and

$$dX_2^{\epsilon}(t) = (a(t)X_2^{\epsilon}(t) + b(t) + \alpha\varphi_t^{\epsilon})dt.$$
(2.6)

Equation (2.5) is a simple stochastic linear equation of Langevin type and its solution is

$$X_1^{\epsilon}(t) = e^{\int_0^t a(u)du} (X_1^{\epsilon}(0) + \sigma \epsilon^{\alpha} \int_0^t e^{-\int_0^s a(u)du} dW_s).$$
(2.7)

And the equation (2.6) is an ordinary differential equation for every fixed ω and its solution is

$$X_{2}^{\epsilon}(t) = e^{\int_{0}^{t} a(u)du} \left[X_{2}^{\epsilon}(0) + \int_{0}^{t} b(s)e^{-\int_{0}^{s} a(u)du}ds + \sigma\alpha \int_{0}^{t} \varphi_{s}^{\epsilon}e^{-\int_{0}^{s} a(u)du}ds \right].$$
(2.8)

Now combining (2.4), (2.7) and (2.8) and noticing that $\alpha \varphi_s^{\epsilon} ds + \epsilon^{\alpha} dW_s = dB_s^{H,\epsilon}$ we can write the approximate signal X_t^{ϵ} in the form

$$X_{t}^{\epsilon} = X_{1}^{\epsilon}(t) + X_{2}^{\epsilon}(t)$$

= $e^{\int_{0}^{t} a(u)du} \left[X_{0} + \int_{0}^{t} b(s)e^{-\int_{0}^{s} a(u)du}ds + \sigma \int_{0}^{t} e^{-\int_{0}^{s} a(u)du}dB_{s}^{H,\epsilon} \right], \quad (2.9)$

where X_0 is assumed a random variable such that $E|X_0|^2 < \infty$.

3 Convergence to the exact solution

We can see that the equation (2.1) satisfies all conditions of Theorem of existence and uniqueness for solution of a fractional stochastic differential equation given in [3]. We will prove that the approximate signal X_t^{ϵ} converges to the fractional X_t that is the exact solution of (2.1). Consider two equations

$$dX_t = (a(t)X_t + b(t))dt + \sigma dB_t^H,$$

$$dX_t^{\epsilon} = (a(t)X_t^{\epsilon} + b(t))dt + \sigma dB_t^{H,\epsilon}.$$

3.1 Theorem 3.1

 X_t^{ϵ} converges to X_t in $L^2(\Omega)$ uniformly with respect to $t \in [0,T]$.

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Proof. We have

$$X_t - X_t^{\epsilon} = a(t) \int_0^t (X_s - X_s^{\epsilon}) ds + \sigma (B_t^H - B_t^{H,\epsilon}).$$

Then

$$||X_t - X_t^{\epsilon}|| \le M || \int_0^t (X_s - X_s^{\epsilon}) ds || + \sigma ||B_t - B_t^{H,\epsilon}||, \qquad (3.1)$$

where $\|.\|$ denote for L^2 -norm and $|a(t)| \le M$ for $t \in [0, T]$, M > 0 due to the fact that a(t) is a continuous function.

In account of (1.7) we can see from (3.1) that

$$\|X_t - X_t^{\epsilon}\| \le M \int_0^t \|X_s - X_s^{\epsilon}\| ds + \sigma K(\alpha) \epsilon^{\frac{1}{2} + \alpha}, 0 \le t \le T.$$
(3.2)

Applying the Gronwall's lemma to (3.2) we get

$$\|X_t - X_t^{\epsilon}\| \le \sigma K(\alpha) \epsilon^{\frac{1}{2} + \alpha} e^{-at}$$
(3.3)

and then

$$\sup_{0 \le t \le T} \|X_t - X_t^{\epsilon}\| \le \sigma K(\alpha) \epsilon^{\frac{1}{2} + \alpha} e^{-aT} \text{ for } a > 0$$
$$\sup_{0 \le t \le T} \|X_t - X_t^{\epsilon}\| \le \sigma K(\alpha) \epsilon^{\frac{1}{2} + \alpha} \text{ for } a < 0$$

So $X_t^{\epsilon} \longrightarrow X_t$ in $L^2(\Omega)$ uniformly with respect to $t \in [0, T]$.

3.2 Collorary 3.1

It follows from Theorem 3.1 and the formula (2.9) that the exact signal X_t can be explicitly expressed as

$$X_t = e^{\int_0^t a(u)du} \left(X_0 + \int_0^t b(s)e^{-\int_0^s a(u)du}ds + \sigma \int_0^t e^{-\int_0^s a(u)du}dB_s^H \right).$$
(3.4)

4 Best state estimate for signal X_t

4.1 Approximation for best state estimate

Consider now an approximate model for state estimate of the signal X_t^{ϵ} form the observation Y_t : Signal X_t^{ϵ} :

$$dX_t^{\epsilon} = (a(t)X_t^{\epsilon} + b(t))dt + \sigma dB_t^{H,\epsilon}.$$
(4.1)

Observation Y_t :

$$dY_t = h(X_t^{\epsilon})dt + dV_t. \tag{4.2}$$

The model (4.1) - (4.2) can be rewritten as follows

$$dX_t^{\epsilon} = \left(a(t)X_t^{\epsilon} + b(t) + \alpha\varphi_t^{\epsilon}\right)dt + \sigma\epsilon^{\alpha}dW_t, \qquad (4.3)$$

$$Y_t^{\epsilon} = \int_0^t h(X_s^{\epsilon}) ds + V_t.$$
(4.4)

where W_t and V_t are two independent standard Brownian motion. Let \mathcal{F}_t^Y be the observation σ -algebra, that is the algebra generated by all random variables Y_s for $s \leq t$:

$$\mathcal{F}_t^Y = \sigma(Y_s, 0 \le s \le t).$$

Also, $\mathcal{F}_t^{Y^{\epsilon}}$ is denoted for the approximate abservation σ -algebra: $\mathcal{F}_t^{Y^{\epsilon}} = \sigma(Y_s^{\epsilon}, 0 \leq s \leq t)$ The best state estimation for approximate signal X_t^{ϵ} denoted by $\widehat{X_t^{\epsilon}}$ based on observation information given by \mathcal{F}_t^Y :

$$\widehat{X_t^{\epsilon}} = E\left(X_t^{\epsilon} | \mathcal{F}_t^{Y^{\epsilon}}\right). \tag{4.5}$$

Denote by ν_t the innovation process that is a $\mathcal{F}_t^{Y^{\epsilon}}$ -martingale:

$$\nu_t = Y_t^{\epsilon} - \int_0^t \widehat{h_s^{\epsilon}} ds, \qquad (4.6)$$

where $\widehat{h_s} = \widehat{h(X_s)} = E(h(X_s)|\mathcal{F}_s^Y), \ 0 \le s \le t$ and by H_t^{ϵ} the following expression

$$H_t^{\epsilon} = a(t)X_t^{\epsilon} + b(t) + \alpha\varphi_t^{\epsilon}.$$
(4.7)

Now we are in position to apply the FKK (Fujisaki-Kallianpur-Kunita) (see [11]) equation to $\widehat{X_t^{\epsilon}}$ and we have

Theorem 4.1 The best state estimate $\widehat{X_t^{\epsilon}}$ is given by the following equation

$$\widehat{X}_{t}^{\widehat{\epsilon}} = \widehat{X}_{0}^{\widehat{\epsilon}} + \int_{0}^{t} \widehat{X_{s}^{\widehat{\epsilon}}H_{s}^{\widehat{\epsilon}}} ds + \int_{0}^{t} \left[\widehat{X_{s}^{\widehat{\epsilon}}h_{s}} - \widehat{X_{s}^{\widehat{\epsilon}}}\widehat{h_{s}^{\widehat{\epsilon}}}\right] d\nu_{s},$$
(4.8)

where the notation \wedge stands for the best state estimate.

4.2 Best state estimation for the exact signal X_t

Now we have to find

$$\widehat{X_t} = E(X_t | \mathcal{F}_t^Y), \tag{4.9}$$

where the signal X_t is given by (3.4). Consider the best approximate state $\widehat{X_t^{\epsilon}} = E(X_t^{\epsilon} | \mathcal{F}_t^{Y^{\epsilon}})$. Put $\epsilon = 1/n, n = 1, 2...$ and denote $X^{(n)}$ for X_t^{ϵ} with $\epsilon = 1/n$. T. H. THAO, T. M. TUONG AND T. P. LOC

Then $\widehat{X_t^{\epsilon}} = \widehat{X_t^{(n)}} = E(X_t^{(n)} | \mathcal{F}_t^{(n)})$ where $\mathcal{F}_t^{Y^{\epsilon}} = \mathcal{F}_t^{(n)}$ is the σ -algebra generated by $(X_0, B_s^{(n)}, V_s, s \leq t)$ with

$$B_t^{(n)} = B_t^{H,1/n} = \int_0^t (t - s - 1/n)^\alpha dW_s.$$
(4.10)

By a change of variable we have

$$B_t^{(n)} = \int_0^{t-1/n} (t-u)^\alpha dW_u = B_{t-1/n} \text{ and } B_s^{(n)} = B_{s-1/n}.$$
 (4.11)

Therefore σ -algebras $\mathcal{F}_t^{(n)} = \sigma(X_0, B_{s-1/n}, V_s, s \leq t), n = 1, 2, \ldots$ form an inscreasing filtration and $\mathcal{F}_t^{(n)} \nearrow \mathcal{F}_t^Y$. By applying the elementary inequality

$$|a+b|^2 \le \frac{1}{2}|a|^2 + \frac{1}{2}|b|^2$$

we can see

$$E|E(X^{(n)}|\mathcal{F}_{t}^{(n)}) - E(X_{t}|\mathcal{F}_{t}^{Y})|^{2} \leq \frac{1}{2}E|E(X_{t}^{(n)} - X_{t}|\mathcal{F}_{t}^{(n)})|^{2} + \frac{1}{2}E|E(X_{t}|\mathcal{F}_{t}^{(n)}) - E(X_{t}|\mathcal{F}_{t}^{Y})|^{2} \leq \frac{1}{2}E|X_{t}^{(n)} - X_{t}|^{2} + \frac{1}{2}E|E(X_{t}|\mathcal{F}_{t}^{(n)}) - E(X_{t}|\mathcal{F}_{t}^{Y})|^{2}. \quad (4.12)$$

In the last side of (4.12) we see that when $n \to \infty$ the first term tends to 0 by Theorem 3.1 for $\epsilon = 1/n$ and the second term converges to 0 as well because of a Levy theorem of convergence of conditional expectation. Finally we can state

Theorem 4.2: $\widehat{X_t}$ can be considered as $L^2 - \lim of \widehat{X_t^{(n)}}$ when $n \to \infty$

$$\widehat{X}_t = L^2 - \lim_{n \to \infty} E(X_t^{(n)}) |\mathcal{F}_t^{(n)}.$$
(4.13)

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